#### Genome Assembly: Lander-Waterman Statistics

#### Sorin Istrail

Center for Computational Molecular Biology Department of Computer Science Brown University, Providence sorin@cs.brown.edu

May 2, 2011

4 3 5 4

- The Lander-Waterman formulas<sup>1</sup> provide a statistical framework for estimating sequencing parameters needed to achieve particular levels of quality for large-scale sequencing projects.
- Suppose we have a long piece of DNA we want to sequence.
- We model the sequencing of fragments by breaking the DNA uniformly at random and then sequencing the DNA at the start of that break.
- Assumption 1: DNA may be broken uniformly at random and all DNA bases are able to be sampled.

<sup>1</sup>E. S. Lander and M. S. Waterman. Genomic mapping by fingerprinting random clones: a mathematical analysis. Genomics, 2(3):231-239, April 1988

We define several parameters common to all sequencing projects for our calculations

- L = read length
- N = number fragments each of length L
- G =original sequence's length (genome length)

• 
$$a = \frac{NL}{G}$$
 or the coverage

# Lander-Waterman Statistics: The three fundamental questions

- Q1. What is the mean portion of the genome covered by contigs?
- Q2. What is the mean number of contigs?
- Q3. What is the mean contig size?

4 B K 4 B K

We must make simplifications via assumptions because the model becomes too complex to analyze analytically. We define now our assumptions.

Assumption 1: DNA may be broken uniformly at random and all DNA region are able to be sampled.

Assumption 2: We assume G >> L and therefore we can safely ignore end effects in our calculations (e.g. reads starting in [G-L+1,G]).

Assumption 3: No repeats on the target genome.

Assumption 4: No errors

## Q1. What is the mean portion of the genome covered by contigs?

- The mean portion of the genome covered by one or more fragments is the probability that a point chosen at random is covered by at least one fragment.
- We can model with a Poisson model.
- Let Y = number of reads whose leftmost end point is located within the interval of length L. This follows a Poisson distribution with mean a.

- The fragments are taken at random from the original full-length sequence so, if end effects are ignored, the left-hand ends of the fragments are independently distributed with a common *uniform distribution* over [0, *G*].
- This implies that any such left-hand end falls in an internal (x, x + h) with probability <sup>h</sup>/<sub>G</sub> and that the number of fragments whose left-hand end falls in this interval has a binomial distribution with mean <sup>NL</sup>/<sub>G</sub>.
- If *N* is large and *h* is small, this distribution is approximately Poisson with mean  $a = \frac{NL}{G}$ .

Image: A Image: A

• The number Y of fragments whose left-end is located within an interval of length L to the left of a randomly chosen point, therefore, has a Poisson distribution with mean a, so that the possibility that at least one fragment arises in this interval is  $1 - Prob(Y = 0) = 1 - e^{-a}$ .

$$P(Y = 0) = e^{-a} \frac{a^0}{0!} = e^{-a}$$
$$P(Y = k) = e^{-a} \frac{a^k}{k!}$$

• *P*(*Y*) together with other properties of homogeneous Poisson processes is enough to provide the answer to these basic questions.

• • = • • = •

### $A1: P(Y \ge 1) = 1 - e^{-a}$

Genome covered by the reads means that every base on the genome is covered by at least one read.

( ) < ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) <

Example. Consider a Shotgun sequencing project. We have a = 8,  $G = 10^4$ , L = 500, N = 1600. Randomly pick the start position of the read (left most position). We have 1600 randomly picked points where reads start. Using  $\frac{NL}{G}$  for the mean for the Poisson we can find the answer to Q1.

	2	4	6	8	10	12
A1: mean	0.86	0.98	0.997	0.9996	0.99995	0.999994
portion of						
the genome						
covered						

### Q2. What is the mean number of contigs?

- Each contig has a *unique rightmost fragment* so that the formula *np* for the mean of the binomial distribution given below shows that the mean number of contigs is the number *N* of fragments multiplied by the probability that a fragment is the rightmost member of a contig.
- Mark all the rightmost fragments.
- Their number equals the number of contigs.
- The probability that no other fragment has its left-hand end point on the fragment in question is  $e^{-a}$ .

• 
$$A2 = Ne^{-a} = Ne^{\frac{-NL}{G}}$$
 is the mean number of contigs.

Example. Recall that  $A2 = Ne^{-a}$ . Let G = 100000 and L = 500. Then

	0.5	0.75	1	1.5	2	3	4	5	6	7
A2:	60.7	70.8	73.6	66.9	54.1	29.9	14.7	6.7	3.0	1.3
mean										
por-										
tion										
of the										
genome										
COV-										
ered										

#### Q3: What is the average contig size?

$$A3 = \frac{e^a - 1}{\lambda} = L \frac{e^a - 1}{a}$$

where  $\lambda = \frac{N}{G}$ ,  $a = \frac{NL}{G}$ ,  $\lambda L = a$ .

( ) < ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) < )
( ) <

- By the Poisson approximation we create the distribution for reads on a genome.
- The distance between the starting point of one read to the starting point of the next overlapping read follows a geometric distribution which we approximate by an exponential distribution with  $\lambda = \frac{N}{G}$ . See Figure 1.

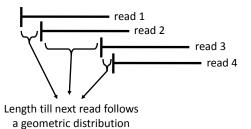


Figure: We approximate the geometric dist by an exponential dist.

- The second fragment will overlap the first one if this distance is less than L in length of the first fragment. This occurs with prob  $1 e^{-a}$
- We then have a series of overlaps until the first failure. Treat overlaps as successes and non-overlaps as failures. The number of successive overlapping fragments before the first non-overlap has a geometric distribution whose mean is e<sup>a</sup> - 1.

- If *n* frags form a contig, the total fragment lengths of the contig is the sum of random distances + L.
- The mean of these random distances is  $\frac{1}{\lambda} \frac{L}{e^a 1}$ . See Figure 2.

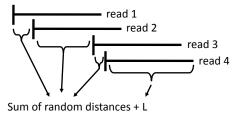
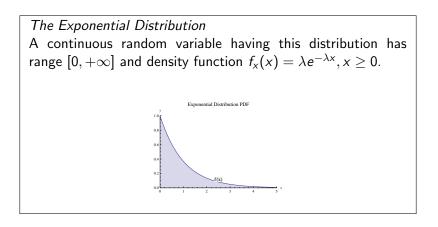


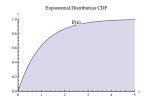
Figure: The contig length is given by the sum of random distances + L.

- The mean of a sum of a random number of random variables show that the mean total of these distances is  $\frac{e^a-1}{\lambda} L$
- Adding *L* from the last fragment we obtain the mean contig size:  $A3 = L\frac{e^a 1}{a}$

- Reads are place uniformly at random along the genome that is the left end point is picked and you read out L bases.
- The mean contig is found by considering the left-hand ends of a succession of fragments starting with the initial left-hand fragment of a given contig.
- The Poisson distribution can approximate the Binomial very well when the number of trials is very large and the probability *p* is very small.
- Using the Poisson approximation for the read distribution, the distance from the left-hand end of this first fragment and the left-hand end of the next fragment has a *geometric* distribution.
- This distribution is closely approximated by the exponential distribution with parameter  $\lambda = \frac{N}{G}$ . For example: N = 1600,  $G = 10^4$ , then  $p = \frac{N}{G} = \frac{1600}{10^4}$ . We make this approximation here.



A single parameter  $\lambda > 0$  characterizes the distribution. The mean is  $\frac{1}{\lambda}$  and the variance is  $\frac{1}{\lambda^2}$ . It has cumulative distribution function  $F_x(x) = \lambda - e^{=\lambda x}, x \ge 0$ 



The Relation of the Exponential to the Geometric Distribution Taking  $\lambda = 1$  and plugging the CDF into the mean of the geometric distribution:  $\frac{p}{1-p} = \frac{1-e^{-a}}{1-(1-e^{-a})} = \frac{1-e^{-a}}{1-(1-e^{-a})} = \frac{1-e^{-a}}{e^{-a}} = \frac{1}{e^{-a}} - 1 = e^a - 1$ 

- Assume the fragments are of constant size *L*.
- The second fragment will overlap the first one if this distance is less than the length *L* of the first fragment.
- This occurs with probability

$$\int_0^L \lambda e^{-\lambda x} dx = 1 - e^{-a}$$

- A further overlap occurs if the next fragment to the right of the second fragment overlaps the second fragment.
- The contig is built this way until such an overlap fails to occur.
- We think of an overlap as "success" and a non-overlap as a "failure."
- The number of successive overlapping fragments, before the first non-overlap has a geometric distribution whose mean is  $\frac{p}{1-p} = e^a 1 \text{ where } p = 1 e^{-a}.$
- If n fragments from a contig, total length of the contig is the length L of the final fragment together with the sum of the n-1 random distances between the left-hand end of any given fragment and the left-hand of the next fragment in the contig to the right.

伺 ト く ヨ ト く ヨ ト

The conditional distribution of an exponential random variable X given that  $0 \le X \le L$  is  $\frac{\lambda e^{-\lambda x}}{1 - e^{-\lambda L}}, \qquad 0 \le x \le L$ The mean of this distribution is  $\int_{0}^{L} x \cdot \frac{\lambda e^{-\lambda x}}{1 - e^{-\lambda L}} dx = \frac{1}{\lambda} - \frac{L}{e^{\lambda L - 1}}$ 

(\* ) \* ) \* ) \* )

The mean of these random distances is  $\frac{1}{\lambda} - \frac{L}{e^a - 1}$  where  $\lambda = \frac{N}{G}$  and  $a = \frac{NL}{G}$ 

The Exponential Distribution vs. Random n for sum of n Random variables The  $X_i$  are independently and identically distributed random variables with mean  $\mu$ . Let  $S_n = X_1 + ... + X_n$ with mean of  $S_n = n\mu$ . However this is assuming n is fixed. We need n to also be a random variable which we will denote by  $\mathcal{N}$ . Assume  $\mathcal{N}$  is independent from  $X_1, ..., X_n$ . The sum is now denoted S. The mean of S is  $E(S) = E(\mathcal{N})E(X)$ 

- So for us the mean of the sum of a random number of random variables is according to above E(S) formula.
- The mean of the random distances is  $\frac{1}{\lambda} \frac{L}{e^a 1}$  where  $\lambda = \frac{N}{G}$
- The mean of the sum of a random variables of random mean is  $\frac{e^a 1}{\lambda} L$ .
- Upon adding length L for the last contig we get

$$A3 = \frac{e^a - 1}{\lambda} = L\frac{e^a - 1}{a}$$

where  $\lambda = \frac{N}{G}$ ,  $a = \frac{NL}{G}$ ,  $\lambda L = a$ .

Example with L = 500.

а	2	4	6	8	10
mean contig size	1,600	6,700	33,500	186,000	1,100,000

Table: An example of the mean contig size calculation for a fixed cover and read length.

(\* ) \* ) \* ) \* )